Some L in it Theorems for Kernel-Smooth Quantile Estimators

Zhou Yong
(Inst of Appl Math , Academia Sinica, Beijing 100080)

Abstract Weak convergence and strong consistency of the remainder term in the Bahadur representation of the sample p - quantile are established. From the results we obtain asymptotic normality and the laws of iterated logarithm for smooth quantile estimator.

Keywords Bahadur representation, sample quantile, empirical process and quantile process **Classification** AM S (1991) 62E20, 62G30/CCL O 212 2, O 212 7

1 Introduction

One characteristic of the distribution that is of interest is the quantile function, which is useful in reliability and medical studies

For the distribution function F, the quantile function is defined by

$$Q(p) = \inf\{x: F(x) \ge p\}, \quad 0$$

A natural estimator of the quantile function Q(p) is the sample quantile function Q_n defined by

$$Q_n(p) = F_n^{-1}(p) = \inf\{x: F_n(x) \ge p\},$$

where $F_n((\bullet))$ is the empirical distribution function (d f).

There are several nonparametric estimators of Q(p) in the literature. For example, the sample quantile function, $F_n^{-1}(p) = \inf\{x: F_n(x) \ge p\}$, $0 \le p \le 1$ has been studied, where $F_n(x)$ is the empirical distribution function based on the sample drawing from popular distribution function F. [2] gave many of the known results concerning F_n^{-1} . A nother approach has been to solve $\widetilde{F}_n(x_p) = p$ for x_p where $\widetilde{F}_n(x) = \int_{-x}^{x} f_n(t) dt$ with $f_n(t)$ being a kernel estimator (see [7]). [9] studied a kernel-type estimator which is the smoothed sample quantile function $F_n^{-1}(p)$ based on the kernel method

The quantile function of the empirical distribution is a step function with jumps corresponding to the observations. The purpose of this paper is to present a smoothed nonpara-

^{*} Received Jan 14, 1995. This work was supported by the postdoctoral grant of China

metric estimator for the quantile function based on the kernel method and obtain some asymptotic results for the smoothed quantile estimator. From these results, we can establish Bahadur representation for this smooth quantile estimator with exact convergent order and exact constant in the order. [9] showed that under general conditions this estimator is strongly consistent, and it performs better than the sample quantile function in the sense of smaller mean squared error, particularly when the size of sample is small

Now, for 0 , define the kernel-type quantile function estimator

$$\widetilde{Q}_{n}(p) = h_{n}^{-1} \int_{0}^{1} Q_{n}(t) k \left[\frac{t-p}{h_{n}} \right], dt = h_{n}^{-1} X_{(i)} \int_{(i-1)/n}^{i/n} k \left[\frac{t-p}{h_{n}} \right] dt, \qquad (1.1)$$

where $X_{(1)}$, $X_{(2)}$, ..., $X_{(n)}$, (i = 1, 2, ..., n - 1) are the order statistics of X_1 , X_2 , ..., X_n and k(t) is a probability density function and $h = h_n$ is a sequence of band width

Let U_i be i i d uniform (0, 1) random variables, and the uniform empirical distribution based on these reduced rv s is then given by

$$G_n(y) = \frac{1}{n} \int_{i=1}^n I(U_i \leq y), \quad y \quad (0,1).$$

where $I(\bullet)$ is the indicator of (\bullet) . Let $e_n(t)$ denote the corresponding uniform process

$$e_n(y) = n^{1/2}(G_n(y) - y), \quad 0 < y < 1.$$

Define q_n , the inverse of G_n , and the uniform empirical process u_n by

$$q_n(y) = \inf\{t: G_n(t) \ge y\},\$$

 $u_n(y) = n^{1/2}(y - a_n(y)), \quad 0 < y < 1.$

In this paper we shall consider the smoothed quantile process

$$\widetilde{\beta}_n(p) = \sqrt{n} f(Q(p)) (\widetilde{Q}_n(p) - Q(p))$$

and the smooth Bahadur- Kiefer process

$$R_n(p) = e_n(Q(p)) + \widetilde{\beta}(p), \quad 0 (1.2)$$

[1] was the first to investigate the distance between the empirical and quantile processes in the case the sample is coming from the uniform U(0, 1) distribution. The best result concerning this problem, is due to [6], he proved the sharpest order in this distance. In this paper we consider the distance between the empirical and smooth quantile processes for the sample coming from general d f

Our main results are the following theorems

Theorem 1 Suppose that F is w ice differentiable on the neighborhood of Q(p), f is continuous and positive near Q(p) and f is continuous on Q(p). Let k be a probability density function w ith finite support (-c,c) for some c>0 and -tk(t)dt=0. Let $\{h=h_n,n\geq 1\}$ be a sequence of bandw id the satisfying

$$\sqrt{n} \ h \log h^{-1} \rightarrow 0 \quad \text{and} \quad \frac{nh}{\log h^{-1}} \rightarrow$$
 (1. 3)

 $as n \rightarrow . Then$

$$\lim P\{n^{\frac{1}{4}}f(Q(p))\widetilde{R}_{n}(p) \leq t\} = 2 \Phi(\frac{t}{u^{\frac{1}{2}}})N(0, p(1-p)), (du).$$

where $\Phi(\bullet)$ is S tandard normal d.f. and N $(0, \sigma^2)$ denotes the normal distribution with expectation zero and variance σ^2 .

Theorem 2 Suppose that f is continuously differentiable and positive on real line. Let k be a probability density function w ith f inite support (-c,c) for same c>0 and tk(t)dt=0. Let $\{h=h_n, n\geq 1\}$ be a sequence of bandwidths satisfying

$$\frac{\sqrt{n \, h \log h^{-1}}}{(\log \log n)^{3/2}} \to 0 \quad \text{and} \quad \frac{nh}{\log h^{-1}} \to \tag{1.4}$$

 $as n \rightarrow . Then$

$$\lim_{n \to \infty} \sup_{n} n^{1/4} (\log \log n)^{-3/4} \left| \widetilde{R}_{n}(p) \right| = 2^{5/4} 3^{-3/4} (p(1-p))^{\frac{1}{4}} \quad a. s \tag{1.5}$$

f or p = (0, 1) f ix ed.

Theorem 3 Suppose that f is continuously differentiable and positive on real line. Let k be a probability density function w ith f in ite support (-c,c) for some c>0 and tk(t), dt=0. Let $\{h=h_n, n\geq 1\}$ be a sequence of bandwidths satisfying

$$\frac{\sqrt{n} h}{(\log h^{-1})^{1/2}} \rightarrow 0 \quad \text{and } \frac{nh}{\log h^{-1}} \rightarrow \tag{1. 6}$$

as $n \rightarrow$

If in addition, that $0 < \inf_{0 < x < 1} f(Q(x)) < .$ Then

$$\lim_{n \to \infty} \sup n^{1/4} (\log n)^{-1/2} (\log \log n)^{-1/4} \sup_{0 \le p \le 1} \left| \widetilde{R}_n(p) \right| = 2^{1/4} \quad \text{a. s.}$$
 (1. 7)

From the law of the iterated logarithm for empirical process we immediately get

Corollary 1 Under all the conditions of Theorem 3 we have

$$\lim_{n \to \infty} \sup \left(\frac{n}{2\log\log n} \right)^{1/2} \sup_{0 \le p \le 1} f\left(Q\left(p \right) \right) \left| \widetilde{Q_n}(p) - Q\left(p \right) \right| = 1 \quad a. s$$

and

$$\lim_{n \to \infty} \inf_{n} \inf_{n} \int_{0}^{\frac{1}{2}} (\log \log n)^{\frac{1}{2}} \sup_{0 \le p \le 1} f(Q(p)) |\widetilde{Q}_{n}(p) - Q(p)| = \pi 8^{-1/2} \quad a. s.$$

2 Proof of the theorems

For convenience of presentation we shall assume throughout that $q_n(y) = 0$ for $y \le 0$ and $u_n(y) = 0$ for $y \ge 1$. We quote some strong approximation results for the empirical process, which are used in the proofs of our theorems

Let $h = h_n$ for simplity.

Lemma 1 If $nh/(\log h^{-1}) \rightarrow as n \rightarrow$, then we have

$$\sup_{0 \le y \le p} \sup_{|t| \le h} |K(y + t, n) - K(y, n)| = O((n h \log h^{-1})^{1/2}) \quad a.s.$$
 (2.1)

where define K(s, t) = 0 as $s \le 0$ and $s \ge 1$.

Proof See [3] Theorem 1. 15. 2

Lemma 2 A ssum e that the regularity conditions in Theorem 1 are satisfied, then we have

$$\sup_{0 \le y \le 1} \sup_{|y| \le h} |u_n(y + s) - u_n(y)| = O((h \log h^{-1})^{1/2}) \ a. \ s$$
 (2.2)

Proof From [3] Theorem 4 5 3, we have

$$n^{-1/2} \sup_{0 < y \le p} \sup_{|s| \le h} |u_n(y + s) - u_n(y)|$$

$$\leq n^{-1/2} \sup_{0 < y \le p} \sup_{|s| \le h} |u_n(y + s) - n^{-1/2}K(y + s, n)|$$

$$+ n^{-1/2} \sup_{0 < y \le p} \sup_{|s| \le h} |K(y + s, n) - K(y, n)|$$

$$+ n^{-1} \sup_{0 < y \le p} \sup_{|s| \le h} |K(y + s, n) - K(y, n)|$$

$$: = A_1 + A_2 + A_3$$
(2.3)

w here

$$A_2 = O(n^{-3/4}(\log n)^{1/2}(\log \log n)^{1/4})$$
 as

For sufficiently large n, we obtain

$$A_1 \le n^{-1/2} \sup_{0 \le y \in n} |u_n(y) - n^{-1/2} K(y, n)| = O(n^{-3/4} (\log n)^{1/2} (\log \log n)^{1/4})$$
 a s

By Lemma 1, we obtain

$$A_3 = O((n^{-1}h\log h^{-1})^{1/2})$$
 a s

Thus the result of Lemma 2 follows from the boundedness of A₁, A₂, A₃

Lemma 3 [[8] **Th** 1A] Suppose that k(t) is a bounded integral function on real line and $\lim_{|t| \to \infty} |tk(t)| = 0$. Define $g_n(x) = \int_{-\infty}^{\infty} g(t) a_n^{-1} k \left(\frac{t-x}{a_n}\right) dt$ Then at every point x of continuity of $g(\bullet)$

$$\lim_{n\to\infty} g_n(x) = g(x) \qquad k(t) dt,$$

as g(x) is uniform continuous, then the equality above holds true uniform by on x.

Proofs of Theorem 1 and Theorem 2 L et $q_n(p) = \inf\{s: G_n(s) \ge p\} = F(Q_n(p))$, and $Q_n(p) = Q(q_n(p))$. For sufficiently large n, we can write

$$\begin{aligned} \left| \widetilde{R}_{n}(p) - R_{n}(p) \right| &= n^{1/2} f(Q(p)) \left| h^{-1} \int_{0}^{1} Q_{n}(t) k \left(\frac{p - t}{h} \right) dt - Q_{n}(p) \right| \\ &= n^{1/2} f(Q(p)) \left| \int_{-c}^{c} \left[Q(q_{n}(t)) - Q(q_{n}(p)) \right] \frac{1}{h} k \left(\frac{p - t}{h} \right) dt \right| \\ &= n^{1/2} f(Q(p)) \left| \int_{-c}^{c} \frac{1}{f(Q(p))} \left[q_{n}(t) - q_{n}(p) \right] \frac{1}{h} k \left(\frac{p - t}{h} \right) dt \right| \\ &+ \int_{-c}^{c} \left[q_{n}(t) - q_{n}(p) \right]^{2} \frac{f(Q(\xi_{t}))}{f^{3}(Q(\xi_{t}))} \frac{1}{h} k \left(\frac{p - t}{h} \right) dt \right| . \end{aligned}$$

$$= I_{1}(p) + I_{2}(p). \tag{2.4}$$

where ξ_t lies between $q_n(p - ht)$ and $q_n(p)$.

By Lemma 2 1 and the fact that yk(y), dy = 0, we have

$$I_{1}(p) = \left| \int_{-c}^{c} (u_{n}(p - ht) - u_{n}(p)) k(t) dt \right| \leq \sup_{|t| \leq c} |u_{n}(p - ht) - u_{n}(p)| \int_{-c}^{c} k(t) dt$$

$$= \sup_{|t| \leq ch} |u_{n}(p + t) - u_{n}(p)| = O((h \log h^{-1})^{1/2}) \quad a.s$$
(2.5)

By the C_r -inequality, we have

$$I_{2}(p) \leq 2n^{-\frac{1}{2}} \left[u_{n}(t) - u_{n}(p) \right]^{2} \frac{\left| f(Q(\xi_{l})) \right|}{f^{3}(Q(\xi_{l}))} \frac{1}{h} k \left(\frac{p-t}{h} \right) dt + 2n^{\frac{1}{2}} \left[t - p \right]^{2} \frac{\left| f(Q(\xi_{l})) \right|}{f^{3}(Q(\xi_{l}))} \frac{1}{h} k \left(\frac{p-t}{h} \right) dt$$
(2.6)

U sing Lemma 1 and Lemma 2 we obtain that the first integration of $(2\ 6)$ is not greater than

$$n^{-\frac{1}{2}} \sup_{0 < t < 1} |u_n(t) - u_n(p)|^2 \quad \frac{|f(Q(\xi_0))|}{|f|^3(Q(\xi_0))} \frac{1}{h} k \left(\frac{p - t}{h}\right) dt$$

$$= O(n^{-\frac{1}{2}} h \log h^{-1}) \quad \text{a s}$$

it then follows from Lemma 2 that

$$\left(\frac{t-p}{h}\right)^{2} \xrightarrow{f(Q(\xi))} \frac{1}{h} k \left(\frac{p-t}{h}\right) dt \rightarrow \frac{f(Q(p))}{f^{3}(Q(p))} t^{2}k(t) dt$$

Hence the third integration of (2 6) is not greater than $O(n^{\frac{1}{2}}h^2)$. Thus we obtain that

$$I_2(p) = O(n^{-1/2}h\log h^{-1} n^{1/2}h^2)$$
 as (2.7)

Since $\sqrt{n} h \log h^{-1} \rightarrow 0 (n \rightarrow n)$, then we have

$$n^{\frac{1}{4}}I_1(p) \rightarrow 0$$
 as and $n^{\frac{1}{4}}I_2(p) \rightarrow 0$ as

By Theorem 1 of [4] we immediately obtain Theorem 1.

Since

$$\frac{\sqrt{n \ h \log n}}{(\log \log n)^{3/2}} \to 0 \quad \text{and} \quad \frac{nh}{\log h^{-1}} \to \dots$$

as $n \rightarrow$, we obtain from (2.5) and (2.7) that

$$n^{1/4} (\log \log n)^{-3/4} I_1 \rightarrow 0$$
 as (2.8)

and

$$n^{1/4}(\log\log n)^{-3/4}I_2 \rightarrow 0$$
 as (2.9)

Thus, it follows from (2 4), (2 5) and (2 7) that

$$n^{1/4} (\log \log n)^{-3/4} |\widetilde{R}_n(p) - R_n(p)| \to 0 \text{ a.s}$$
 (2.10)

Therefore, Theorem 2 follows from (2 10) and [6] Theorem 1.

Proof of Theorem 3W e can prove Theorem 3 by the same argument used in the proof of Theorem 2, we may write

$$\sup_{0 \le p \le 1} \left| \widetilde{R}_{n}(p) - R_{n}(p) \right| \le \sup_{0 \le p \le 1} I_{1}(p) + \sup_{0 \le p \le 1} I_{2}(p). \tag{2} 11$$

By assumption (1. 14) and $0 < \inf_{0 \le x \le 1} f(Q(x)) < \infty$, we have

$$\sup_{0 \le p \le 1} \frac{\left| f\left(Q\left(\xi_{l}\right)\right)\right|}{f^{3}\left(Q\left(\xi_{l}\right)\right)} k\left(t\right) dt \le \sup_{0 \le p \le 1} \frac{\left| f\left(Q\left(p\right)\right)\right|}{f^{3}\left(Q\left(p\right)\right)} < .$$

By Lemma 1, it follows that

$$\sup_{0 \le p \le 1} I_2(p) \le n^{-\frac{1}{2}} \sup_{0 \le p \le 1} \sup_{|t| \le c} |q_n(p - ht) - q_n(p)|^2 f(Q(p)) \xrightarrow{c} \frac{|f(Q(\xi_t))|}{|f|^3 Q(\xi_t)} k(t) dt$$

$$= O(n^{-\frac{1}{2}} h \log h^{-1} - n^{1/2} h^2) \quad \text{a s}$$

By Lemma 1, we obtain

$$\sup_{0 \le p \le 1} I_1(p) \le \sup_{0 \le p \le 1} \sup_{|t| \le ch} \left| u_n(p+t) - u_n(p) \right| = O\left(\left(h \log (n/h) \right)^{1/2} \right) \quad \text{a s}$$

Since

$$\frac{\sqrt{n h \log h^{-1}}}{\log n} \to 0 \quad \text{and} \quad \frac{nh}{\log h^{-1}} \to \quad ,$$

we obtain

$$n^{1/4} (\log n)^{-1/2} (\log \log n)^{-1/4} \sup_{0 \le p \le 1} I_1(p) \to 0$$
 as

and

$$n^{1/4}(\log n)^{-1/2}(\log\log n)^{-1/4}\sup_{0\le p\le 1}I_2(p)\to 0$$
 as

Therefore, we have

$$n^{1/4} (\log n)^{-1/4} (\log \log n)^{-1/4} \sup_{0 \le s \le 1} |\widetilde{R}_n(p) - R_n(p)| \to 0 \text{ a.s.}$$
 (2.12)

[3] Theorem 5. 2. 2, the result of Theorem 3 follows from (2. 12)

References

- [1] R. R. Bahadur, A note on quantiles in large samples, Ann Math Statist, 37(1966), 577-580
- [2] M. Csorgo, Quantile Process and their Application in Statistics, SAM, Philadelphia PA, 1983
- [3] M. Csorgo and P. R & Ész, Strong Approx in ation in Probability and Statistics, Academic Press, 1981.
- [4] M. Falk, Weak convergence of the renainder term in the Bahadur representation of extreme quantile, Statist Probab Lett, 8(1990), 47-50
- [5] J. Kiefer, On B ahadur s representation of sample quantile, Ann Math Statist, 38(1967), 1323-1342
- [6] J. Kiefer, Deviations between the sample quantile and the sample D. F., Nonparametric Techniques in Statist Inference (M. L. Puri, Ed) 299-319, Cambridge Univ. Press London, 1970
- [7] E.A. Nadaraya, Some new estimates for distribution functions, Theory Probab Appl, 9(1964), 497-500
- [8] E. Parzen, On estimation of a probability density function and mode, Ann. Math. Statist, 33 (1962), 1065-1076
- [9] S. S. Yang, A smooth nonparametric estimator of aquantile function, J. Amer. Statist Assoc, 80 (1985), 1004-1011.