

### 线性模型中误差方差估计的指数收敛速度\*

苏 淳

(中国科学技术大学)

考虑线性回归模型

$$y_j = x_j' \beta + \mu_j, \quad j = 1, 2, \dots \quad (1)$$

其中  $\{\mu_j\}$  为独立的试验误差序列, 满足条件:

$$E\mu_j = 0, \text{Var} \mu_j = \sigma^2, \quad 0 < \sigma^2 < \infty, \quad j = 1, 2, \dots \quad (2)$$

根据最小二乘法规定和模型自身特点, 通常, 在考虑极限性质时, 给  $\sigma^2$  以估计.

$$\hat{\sigma}_n^2 = \frac{1}{n-r} \left\{ \sum_{j=1}^n \mu_j^2 - \sum_{k=1}^r \left( \sum_{j=1}^n a_{nk} \mu_j \right)^2 \right\}. \quad (3)$$

其中  $r$  是一个与设计矩阵  $X_n = (x_1 \dots x_n)$  有关的常正整数, 数列  $\{a_{nkj}\}$  满足等式  $\sum_{j=1}^n a_{nkj} a_{nlj} = \delta_{kl}$ , 这里  $\delta_{kl}$  为 Kronecker 符号,  $n = 1, 2, \dots$ . 本文旨在讨论  $\hat{\sigma}_n^2 \xrightarrow{\text{in pr.}} \sigma^2$  的指数收敛速度, 主要结果是:

**定理** 如果  $\mu_1, \mu_2, \dots$  独立, 满足 (2) 式, 对每个  $\eta > 0$ , 存在仅依赖于  $\eta$  的常数  $M_\eta > 0, t_\eta > 0$ , 使得对一切  $|t| \leq t_\eta$ , 一切  $n$ , 有

$$\prod_{j=1}^n E e^{t(\mu_j^2 - \sigma^2)} = E e^{t \sum_{j=1}^n (\mu_j^2 - \sigma^2)} \leq M_\eta e^{\eta |t|} \quad (4)$$

又如果对一切  $n, k, j$ , 存在常数  $0 < L < \infty$ , 使得  $0 \leq a_{nkj}^2 \leq \frac{L}{n}$ . 则对任何充分小的  $\varepsilon > 0$ , 存在仅依赖于  $\varepsilon$  的常数  $A = A(\varepsilon) > 0$ , 和  $0 < \rho = \rho(\varepsilon) < 1$ , 使得

$$P\{|\hat{\sigma}_n^2 - \sigma^2| \geq \varepsilon\} \leq A \rho^n. \quad (5)$$

为证定理, 先证一个引理.

**引理** 设  $\mu_1, \mu_2, \dots$  独立、对称, 满足条件 (2)、(4), 又存在  $0 < L < \infty$ , 使  $0 < a_{nj}^2 \leq \frac{L}{n}$ , 且  $\sum_{j=1}^n a_{nj}^2 = 1$ , 则对任何充分小的  $\varepsilon > 0$  和充分大的  $n$ , 存在仅依赖于  $\varepsilon$  的常数  $A > 0, 0 < \rho < 1$ , 使得

$$P\left\{\left|\left(\sum_{j=1}^n a_{nj} \mu_j\right)^2 - \sigma^2\right| \geq \frac{n-r}{r} \varepsilon\right\} \leq A \rho^n, \quad (6)$$

**证** 易知, 只须证明

$$J_n = P\left\{\left(\sum_{j=1}^n a_{nj} \mu_j\right)^2 \geq \frac{n-r}{r} \varepsilon\right\} \leq A \rho^n. \quad (7)$$

为方便计, 记  $a_{nj} = a_j$ , 并设  $\sigma^2 = \frac{1}{L}$ . 若不然, 可令  $\tilde{\mu}_j = \frac{\mu_j}{\sigma \sqrt{L}}$ , 以  $\varepsilon_1 = \frac{\varepsilon}{\sigma^2 L}$  代  $\varepsilon$ , 由条件

(4) 中  $\eta > 0$  的任意性, 知存在  $\tilde{M}_\eta > 0, \tilde{t}_\eta > 0$ , 使当  $|t| \leq \tilde{t}_\eta$  时对一切  $n$ , 仍有  $\prod_{j=1}^n E e^{t(\tilde{\mu}_j^2 - 1)} = \prod_{j=1}^n E e^{\frac{t}{\sigma^2 L} (\mu_j^2 - \sigma^2)} \leq \tilde{M}_\eta e^{|t| \eta}$ . 事实上, 只要令  $\tilde{\eta} = \sigma^2 L \eta$ , 取  $\tilde{M}_\eta = M \tilde{\eta}$ , 并使  $\left|\frac{t}{\sigma^2 L}\right| \leq |\tilde{t}_\eta|$  即可.

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由 Чебышев 不等式, 对一切  $t > 0$ , 有

$$\begin{aligned} J_n &\leq e^{t(1-\frac{n}{r})} E e^{t(\sum_{j=1}^n a_j \mu_j)^2} = e^{t(1-\frac{n}{r})} \sum_{k=0}^{\infty} \frac{t^k E \left( \sum_{j=1}^n a_j \mu_j \right)^{2k}}{k!} \\ &= e^{t(1-\frac{n}{r})} \left\{ \left( \sum_{0 < k < \frac{\varepsilon}{r} n} + \sum_{\frac{\varepsilon}{r} n < k < \frac{\ln 2}{2} n} + \sum_{k > \frac{\ln 2}{2} n} \right) \frac{t^k E \left( \sum_{j=1}^n a_j \mu_j \right)^{2k}}{k!} \right\} \\ &\triangleq J_{1n} + J_{2n} + J_{3n} \end{aligned} \quad (8)$$

由  $\mu_1, \mu_2, \dots$  独立、对称, 知  $E \left( \sum_{j=1}^n a_j \mu_j \right)^{2k}$  中的非零项皆具形式  $\frac{(2k)!}{(2i_1)! (2i_2)! \dots (2i_l)!} E(a_{j_1} \mu_{j_1})^{2i_1} E(a_{j_2} \mu_{j_2})^{2i_2} \dots E(a_{j_l} \mu_{j_l})^{2i_l}$ , 其中  $i_1 + \dots + i_l = k$ . 由 Stirling 公式知

$$\begin{aligned} \frac{(2k)!}{(2i_1)! (2i_2)! \dots (2i_l)!} &\leq \frac{(2k)^{2k + \frac{1}{2} i_1^{i_1 + \frac{1}{2} i_2^{i_2 + \frac{1}{2} \dots i_l^{i_l + \frac{1}{2}}}} e^{\frac{1}{12} [\frac{1}{2k} + \sum_{j=1}^l \frac{1}{i_j}]}]}{k^{k + \frac{1}{2} (2i_1)^{2i_1 + \frac{1}{2} (2i_2)^{2i_2 + \frac{1}{2} \dots (2i_l)^{2i_l + \frac{1}{2}}}} \\ &\leq \frac{k^k}{i_1^{i_1} i_2^{i_2} \dots i_l^{i_l}} 2^{\frac{1}{2} (1-l)} e^{\frac{1}{12} \min(k, n)} \leq l^k 2^{\frac{1}{2} (1-l)} e^{\frac{1}{12} \min(k, n)} \end{aligned} \quad (9)$$

这里, 不等式  $\frac{k^k}{i_1^{i_1} i_2^{i_2} \dots i_l^{i_l}} \leq l^k$  可由  $f(t) = t \ln t$  在  $(0, +\infty)$  中的凸性得出. 又通过对  $g_1(l) = l^k 2^{\frac{1}{2} (1-l)}$  求导, 知

$$l^k 2^{\frac{1}{2} (1-l)} \leq k^k 2^{\frac{1}{2} (1-k)} \quad (1 \leq l \leq k) \quad (10)$$

综合 (9) (10) 二式, 知当  $\frac{\varepsilon}{r} n < k < \frac{\ln 2}{2} n$  时, 有

$$E \left( \sum_{j=1}^n a_j \mu_j \right)^{2k} \leq \sqrt{2} \left( \frac{\ln 2}{2} \right)^{\frac{\varepsilon}{r} n} \cdot n^k \left( \frac{e^{\frac{1}{12}}}{\sqrt{2}} \right)^k E \left( \sum_{j=1}^n a_j^2 \mu_j^2 \right)^k \leq \sqrt{2} \left( \frac{\ln 2}{2} \right)^{\frac{\varepsilon}{r} n} \cdot L^k E \left( \sum_{j=1}^n \mu_j^2 \right)^k$$

所以

$$J_{2n} \leq \sqrt{2} e^{t(1-\frac{n}{r})} \left( \frac{\ln 2}{2} \right)^{\frac{\varepsilon}{r} n} \sum_{\frac{\varepsilon}{r} n < k < \frac{\ln 2}{2} n} \frac{t^k L^k E \left( \sum_{j=1}^n \mu_j^2 \right)^k}{k!} \leq \sqrt{2} e^{t(1-\frac{n}{r})} e^{t \frac{\varepsilon}{r} n} \left( \frac{\ln 2}{2} \right)^{\frac{\varepsilon}{r} n} E e^{L t \sum_{j=1}^n (\mu_j^2 - \frac{1}{L})}$$

命  $\eta = \frac{1}{L} > 0$ , 取  $0 < t < \frac{t_n}{L}$  足够小, 使  $0 < \rho_1 \triangleq e^{t(2-\frac{\varepsilon}{r})} \left( \frac{\ln 2}{2} \right)^{\frac{\varepsilon}{r}} < 1$ , 由条件 (4) 知: 存在仅依赖于  $\varepsilon$  的常数  $A_1 > 0$ , 使

$$J_{2n} \leq A_1 \rho_1^n \quad (11)$$

又通过对  $g_2(l) = \left( \frac{1}{n} \right)^{\frac{l-2}{2} n} 2^{\frac{1}{2} (n-l)}$  取导数, 可以证得: 当  $k \geq \frac{\ln 2}{2} n$  且  $1 \leq l \leq n$  时, 有

$$\left( \frac{1}{n} \right)^k 2^{\frac{n-l}{2}} \leq 1 \text{ 记 } \rho_2 = \frac{e^{\frac{1}{12}}}{\sqrt{2}} < 1, \text{ 我们就有}$$

$$l^k 2^{\frac{1}{2} (1-l)} e^{\frac{1}{12} \min(k, n)} = \sqrt{2} n^k 2^{-\frac{n}{2}} e^{\frac{1}{12} \min(k, n)} \cdot \left( \frac{1}{n} \right)^k 2^{\frac{1}{2} (n-l)} \leq \sqrt{2} n^k \rho_2^n.$$

结合 (9) 式, 立刻得到

$$J_{3n} \leq e^{t(1-\frac{n}{r})} \sqrt{2} \rho_2^n \sum_{k > \frac{\ln 2}{2} n} \frac{(nt)^k E \left( \sum_{j=1}^n a_j^2 \mu_j^2 \right)^k}{k!} \leq$$

$$e^{t(1-\frac{n}{r})} \sqrt{2} \rho_2^n \sum_{k > \frac{t_0 n}{2}} \frac{(Lt)^k E \left( \sum_{j=1}^n \mu_j^2 \right)^k}{k!} \leq \sqrt{2} \rho_2^n e^{t(1-\frac{n}{r})} e^{t_0 n} E e^{Lt \sum_{j=1}^n (\mu_j^2 - \frac{1}{2})}$$

仍命  $\eta = \frac{1}{L} > 0$ , 取  $0 < t < \frac{t_0}{L}$  足够小, 使  $0 < \rho_3 \triangleq e^{t(2-\frac{\varepsilon}{r})} \rho_2 < 1$ , 由条件 (4) 得

$$J_{3n} \leq \sqrt{2} M_\eta \rho_2^n e^{t(1-\frac{n}{r})} e^{t_0 n} \triangleq A_2 \rho_3^n \quad (12)$$

取定  $0 < t_0 < \frac{t_0}{L}$ , 使其满足  $J_{2n}$ ,  $J_{3n}$  中的要求, 我们来估计  $J_{1n}$ . 取  $\theta$ :  $0 < \theta < 1 - \frac{e^{\frac{1}{2}}}{\sqrt{2}}$ , 取

$\eta_1 = \frac{1}{L} \theta t_0$ , 由条件 (4) 知: 存在  $M_{\eta_1} > 0$ ,  $t_{\eta_1} > 0$ , 使当

$$0 < \frac{\varepsilon L}{r} < t_{\eta_1} = t_{\eta_1}^t \quad (13)$$

时, 就有

$$E e^{\frac{\varepsilon L}{r} \left( \sum_{j=1}^n (\mu_j^2 - \frac{1}{2}) \right)} \leq M_{\eta_1} e^{\frac{\varepsilon L}{r} \eta_1 n} = M_\theta e^{t_0 n} \quad (14)$$

由于可取  $\varepsilon > 0$  充分小, 所以 (13) (14) 的成立是没有问题的. 不难证明:

$$\lim_{n \rightarrow \infty} e^{-\frac{\varepsilon}{r} n} \sum_{k > [\frac{\varepsilon}{r} n] + 1} \frac{\left( \frac{\varepsilon}{r} n \right)^k}{k!} = \frac{1}{2}$$

所以当  $n$  充分大, 就有  $\sum_{k > \frac{\varepsilon}{r} n} \frac{\left( \frac{\varepsilon}{r} n \right)^k}{k!} \geq \frac{1}{4} e^{\frac{\varepsilon}{r} n}$ . 于是

$$\begin{aligned} E \left( \sum_{j=1}^n a_j^2 \mu_j^2 \right)^{[\frac{\varepsilon}{r} n]} \frac{e^{\frac{\varepsilon}{r} n}}{4} &\leq E \left( \sum_{j=1}^n a_j^2 \mu_j^2 \right)^{[\frac{\varepsilon}{r} n]} \sum_{k > \frac{\varepsilon}{r} n} \frac{\left( \frac{\varepsilon}{r} n \right)^k}{k!} \\ &\leq \sum_{k > \frac{\varepsilon}{r} n} \frac{\left( \frac{\varepsilon}{r} n \right)^k}{k!} \left\{ 1 + E \left( \sum_{j=1}^n a_j^2 \mu_j^2 \right)^k \right\} \leq e^{\frac{\varepsilon}{r} n} + E e^{\frac{\varepsilon L}{r} \left( \sum_{j=1}^n \mu_j^2 \right)}, \end{aligned}$$

从而  $E \left( \sum_{j=1}^n a_j^2 \mu_j^2 \right)^{[\frac{\varepsilon}{r} n]} \leq 4 + 4 E e^{\frac{\varepsilon L}{r} \left( \sum_{j=1}^n (\mu_j^2 - \frac{1}{2}) \right)} \leq 4 + 4 M_\theta e^{t_0 n}$ . (15)

于是, 当  $1 \leq k \leq \frac{\varepsilon}{r} n$  时, 由 (9) (10) 可知

$$\begin{aligned} E \left( \sum_{j=1}^n a_j \mu_j \right)^{2k} &\leq e^{\frac{k}{12}} k^k 2^{\frac{1}{2}(1-k)} E \left( \sum_{j=1}^n a_j^2 \mu_j^2 \right)^k \leq \sqrt{2} \left( \frac{e^{\frac{1}{2}} \varepsilon}{\sqrt{2} r} \right)^n n^k E \left( \sum_{j=1}^n a_j^2 \mu_j^2 \right)^k \\ &\leq \sqrt{2} \left( \frac{e^{\frac{1}{2}} \varepsilon}{\sqrt{2} r} \right)^n n^k \left\{ 1 + E \left( \sum_{j=1}^n a_j^2 \mu_j^2 \right)^{[\frac{\varepsilon}{r} n]} \right\} \leq \sqrt{2} \left( \frac{e^{\frac{1}{2}} \varepsilon}{\sqrt{2} r} \right)^n n^k \left\{ 5 + 4 M_\theta e^{t_0 n} \right\}, \end{aligned}$$

所以

$$\begin{aligned} J_{1n} &\leq \sqrt{2} e^{t(1-\frac{n}{r})} \left\{ 5 + 4 M_\theta e^{t_0 n} \right\} \sum_{k=0}^{[\frac{\varepsilon}{r} n]} \frac{1}{k!} \left( \frac{e^{\frac{1}{2}} \varepsilon n t}{\sqrt{2} r} \right)^k \\ &\leq \sqrt{2} \left\{ 5 + 4 M_\theta e^{t_0 n} \right\} e^{t(1-\frac{n}{r})} e^{t_0 n} \frac{e^{\frac{\varepsilon n t}{\sqrt{2} r}}}{\sqrt{2} r}. \end{aligned}$$

取  $t = t_0$ , 注意到  $\theta$  的取法, 立知  $J_{1n} \leq A_3 \rho_1^n$ . 由此式结合 (11) (12) 知引理获证. 下面来

证定理本身.

众所周知,若随机变量  $X$  有有限方差,则其中位数  $m(X)$  满足  $E(X) - \sqrt{2\text{Var}(X)} \leq m(X) \leq E(X) + \sqrt{2\text{Var}(X)}$ . 我们有

$$\left. \begin{aligned} E\left(\sum_{j=1}^n a_{nkj}\mu_j\right) &= \sum_{j=1}^n a_{nkj}E\mu_j = 0 \\ \text{Var}\left(\sum_{j=1}^n a_{nkj}\mu_j\right) &= \sum_{j=1}^n a_{nkj}^2 \text{Var}(\mu_j) = \sigma^2 \end{aligned} \right\} \text{对一切 } n, k,$$

故有  $\left| m\left(\sum_{j=1}^n a_{nkj}\mu_j\right) \right| \leq \sqrt{2}\sigma$ . 记  $\mu_j$  的对称化随机变量为  $\tilde{\mu}_j$ , 由引理及弱对称不等式知:

$$P\left\{\left|\sum_{j=1}^n a_{nkj}\mu_j - m\left(\sum_{j=1}^n a_{nkj}\mu_j\right)\right| \geq \sqrt{\sigma^2 + \frac{n-r}{r}}\varepsilon\right\} \leq 2P\left\{\left|\sum_{j=1}^n a_{nkj}\tilde{\mu}_j\right| \geq \sqrt{\sigma^2 + \frac{n-r}{r}}\varepsilon\right\} \leq A\rho^n.$$

注意到  $m\left(\sum_{j=1}^n a_{nkj}\mu_j\right) \leq \sqrt{2}\sigma$ , 即知  $P\left\{\left|\sum_{j=1}^n a_{nkj}\mu_j\right| \geq \sqrt{\sigma^2 + \frac{n-r}{r}}\varepsilon\right\} \leq A\rho^n$ . (16)

我们有  $\hat{\sigma}_n^2 - \sigma^2 = \frac{1}{n-r}\left\{\sum_{j=1}^n (\mu_j^2 - \sigma^2) - \sum_{k=1}^r \left[\left(\sum_{j=1}^n a_{nkj}\mu_j\right)^2 - \sigma^2\right]\right\}$ . 由(16)式可知

$$\begin{aligned} P\left\{\frac{1}{n-r}\left|\sum_{k=1}^r \left[\left(\sum_{j=1}^n a_{nkj}\mu_j\right)^2 - \sigma^2\right]\right| \geq \varepsilon\right\} \\ \leq \sum_{k=1}^r P\left\{\left|\left(\sum_{j=1}^n a_{nkj}\mu_j\right)^2 - \sigma^2\right| \geq \frac{n-r}{r}\varepsilon\right\} \leq A\rho^n. \end{aligned} \quad (17)$$

取  $0 < \eta < \varepsilon$ ,  $0 < t < t_n$ , 由条件(4)知

$$P\left\{\frac{1}{n-r}\sum_{j=1}^n (\mu_j^2 - \sigma^2) \geq \varepsilon\right\} \leq e^{-(n-r)\varepsilon t} E e^{t \sum_{j=1}^n (\mu_j^2 - \sigma^2)} \leq M_n e^{r\varepsilon t} e^{-t(\eta-\varepsilon)n} \leq c_1 \tilde{\rho}_1^n. \quad (18)$$

同理可知

$$P\left\{\frac{1}{n-r}\sum_{j=1}^n (\mu_j^2 - \sigma^2) \leq -\varepsilon\right\} \leq c_2 \tilde{\rho}_2^n. \quad (19)$$

综合(17)–(19)知定理获证.

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#### 参 考 文 献

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## Exponential Convergence Rates of Error-variance Estimates in Linear Models

Sn Chun

### Abstract

Suppose given a linear model  $y_j = x_j' \beta + \mu_j$ ,  $j = 1, 2, \dots$ . The random errors all have a mean zero and unknown variance  $\sigma^2$ ,  $0 < \sigma^2 < \infty$ . Let  $\hat{\sigma}_n^2$  be the estimate of  $\sigma^2$  based on the residual sum of squares and calculated from  $(x_j, y_j)$ ,  $j = 1, \dots, n$ . In this paper we show that if  $\mu_1, \mu_2, \dots$ , are independent but not necessarily identically distributed, and some further conditions on  $\{\mu_j\}$  and  $(x_1 | \dots | x_n)$  are satisfied, then for any  $\varepsilon > 0$  there exist constant  $\rho_\varepsilon$ ,  $0 < \rho_\varepsilon < 1$ , Such that

$$P(|\hat{\sigma}_n^2 - \sigma^2| \geq \varepsilon) = O(\rho_\varepsilon^n).$$